

Integrating regional climate change into allometric equations for estimating tree aboveground biomass of Masson pine in China

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Received: 14 September 2016 / Accepted: 11 April 2017 / Published online: 15 May 2017
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Abstract

• **Key message** A climate-sensitive aboveground biomass (AGB) equation, in combination with nonlinear mixed-effects modeling and dummy variable approach, was developed to examine how climate change may affect the allometric relationships between tree diameter and biomass. We showed that such changes in allometry need to be taken into account for estimating tree AGB in Masson pine.

• **Context** As a native species and being widely distributed in subtropical China, Masson pine (*Pinus massoniana* Lamb.) forests play a pivotal role in maintaining forest ecosystem functions and mitigation of carbon concentration increase at

the atmosphere. Traditional biomass allometric equations do not account for a potential effect of climate on the diameter–biomass relationships. The amplitude of such an effect remains poorly documented.

• **Aims** We presented a novel method for detecting the long-term (2041–2080) effects of climate change on the diameter–biomass relationships and the potential consequences for long-term changes of biomass accumulation for Masson pine.

• **Methods** Our approach was based on a climate-sensitive AGB model developed using a combined nonlinear mixed-effects model and dummy variable approach. Various climate-related variables were evaluated for their contributions to model improvement. Heteroscedasticity was accounted for by three residual variance functions: exponential function, power function, and constant plus function.

• **Results** The results showed that diameter at breast height, together with the long-term average of growing season temperature, total growing season precipitation, mean temperature of wettest quarter, and precipitation of wettest quarter, had significant effects on values of AGB. Excessive rain during the growing season and high mean temperature in the wettest quarter reduced the AGB, while a warm growing season and abundant precipitation in the wettest quarter increased the AGB.

Handling Editor: Shuqing Zhao

Contribution of the co-authors L. Fu was the principal investigator for this study. He undertook the main data analyses and wrote the manuscript. X. Lei supervised the research project and contributed to interpretation of results and the writing and revision of earlier drafts until final approval. W. Zeng, S. Tang, P. Marshall, L. Cao, X. Song, and J. Liang contributed to the presentation of the statistical results and interpretations of the results and revised the text. L. Yu and Z. Hu. contributed to climate data preparation.

Electronic supplementary material The online version of this article (doi:10.1007/s13595-017-0636-z) contains supplementary material, which is available to authorized users.

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• **Conclusion** Climate change significantly affected the allometric scale of biomass equation. The new climate-sensitive allometric model developed in this study may improve biomass predictions compared with the traditional model without climate effects. Our findings suggested that the AGB of Masson pine trees with the same diameter at breast height under three climate scenarios including representative concentration pathway (RCP) 2.6, RCP 4.5, and RCP 8.5 in the future period 2041–2080 would increase by $24.8 \pm 32.7\%$ (mean \pm standard deviation), $27.0 \pm 33.4\%$, and $27.7 \pm 33.8\%$ compared with the constant climate (1950–2000), respectively. As a consequence, we may expect a significant regional variability and uncertainty in biomass estimates under climate change.

Keywords Masson pine · Subtropical China · Climate change · Climate-sensitive aboveground biomass model · Nonlinear mixed-effects model · Dummy variable approach

1 Introduction

Global climate change affects all aspects of forest community structure and ecosystem production, such as seedling establishment (Daniels and Veblen 2004), tree growth and mortality (Peng et al. 2011; Subedi and Sharma 2013; Lei et al. 2016), tree species distribution (Hamann and Wang 2006; Sinervo et al. 2010), competition among species (Hamann and Wang 2006; Boisvert-Marsh et al. 2014; Monleon and Lintz 2015), and forest biomass and productivity (Jiang et al. 2015; Tian et al. 2016). Forests play an important role in mitigation of climate change by removing carbon from the atmosphere and storing it in biomass. Consequently, it becomes increasingly important to understand how climate affects forest biomass. The information on the effect of climate on forest biomass can provide valuable insights for formulating forest management strategies, with respect to sustainability and climate change adaptations (Marlon et al. 2008; Lei et al. 2016).

Process-based models have been used to describe climatic effects on forest biomass and productivity for various climatic zones and tree species (Eggers et al. 2008; Shuman and Shugart 2009; Poudel et al. 2011). Not surprisingly, the responses are inconsistent. For the **Mediterranean climate** zone in Europe, it has been speculated that climate change may increase forest productivity by 12–14% and carbon stock by 23–31% in the next 50 years (Eggers et al. 2008). For the Eurasian boreal forests, Shuman and Shugart (2009) found that a significant relationship existed between altered precipitation and biomass. Their analysis also showed that a modest increase in temperature of 2 °C over 200 years would have no

significant effects on biomass. For the temperate continental climate zone in north-central Sweden, an average regional temperature increase of 4 °C over the next 100 years was suggested to increase annual forest productivity by 33% and the potential annual harvest by 32% (Poudel et al. 2011). The temperate moist forests with cool temperatures and moderately high precipitation were found to have the highest biomass and carbon stocks (Keith et al. 2009). For the tropical forests, Stegen et al. (2011) found that aboveground biomass was inversely related to mean annual temperature in wet tropical forest sites and positively related in moist tropical forests and that biomass was positively related to annual precipitation in dry tropical forests. In addition, empirical studies have revealed that rising temperatures increase the availability of soil nitrogen and when combined with a longer growing season, it is expected to increase overall tree growth/biomass (Strömgren and Linder 2002). However, on some sites, increasing temperature may result in temperature-induced drought stress and thus reduce tree growth/biomass (Wilmking et al. 2004). To our knowledge, the magnitude of climate effects on forest biomass remains poorly characterized for the subtropical forests. Traditional biomass allometric equations do not account for a potential climate–biomass relationship and may produce bias for biomass and carbon estimation (Xiang et al. 2011; Zeng et al. 2011; Fu et al. 2016).

The subtropical climate zone in China is located in the eastern part of the country between latitudes 22° and 34° N and longitudes 98° and 123° E (Yang et al. 2006). Its climate is characterized by hot, humid summers and chilly winters (Yang et al. 2006). The subtropical area in China is stratified into three different secondary climate zones (north-, middle-, and south-subtropical zones) based on the mean air temperature and the number of days without frost (Huang 1992). This zone accounts for 63% of the total plantation area (SFA 2012) and about 25% of the total forest biomass in China (Lin et al. 2012). China's subtropical forests, which include Masson pine (*Pinus massoniana* Lamb.) and Chinese fir (*Cunninghamia lanceolata* Lamb. Hook) dominated stands, supply about two thirds of the harvested roundwoods, and are crucial to the economic and social development of rural populations, as well as regional carbon storage and cycling (SFA 2012). Masson pine dominates the regional forests and plays an important role in maintaining the structure and functions of the forest ecosystems in the entire subtropical zone. Masson pine covers a total of 1.13 million ha of forested land, spread throughout 17 provinces in China. Owing to its wide distribution and ability to grow in poor site conditions and regenerate naturally, as well

as its rapid growth rate, Masson pine can greatly contribute to carbon sequestration and storage. Several individual tree biomass models have been developed to estimate the biomass and carbon storage of Masson pine in the subtropical China (Xiang et al. 2011; Zeng et al. 2011; Fu et al. 2016); however, these models did not consider the effects of climate change on biomass allometry.

The limited knowledge of how changes in the regional climate affect the biomass of Masson pine in the subtropical China makes it challenging to prepare for the potential consequences of climate change. According to the Abdus Salam International Centre for Theoretical Physics (ICTP) Regional Climate Model version 3 (RegCM3) with the A1B scenario (IPPC 2001), the mean temperatures in the periods of December–January–February (DJF) and June–July–August (JJA) in China are expected to rise by 5.5 and 5.0 °C, respectively, by the end of the twenty-first century (Gao et al. 2012). Therefore, the objectives of this study were (1) to develop an individual climate-sensitive aboveground biomass (AGB) model with Masson pine as a case study and (2) to examine how climate change may affect the allometric relationships between diameter and biomass and assess the potential error in predicting AGB of Masson pine under different climate scenarios based on the developed AGB model.

2 Materials and methods

2.1 Biomass data

The biomass data used in this study were collected according to the protocol of data collection for tree biomass modeling which was drafted by the State Forestry Administration of China (SFA 2014). A total of 150 Masson pine trees from both natural and planted forests in subtropical China were destructively sampled in 2009 (Fig. 1). The sample trees were located in a broad area in which various forests of Masson pine were distributed and covered nine provinces (Jiangsu, Zhejiang, Anhui, Fujian, Jiangxi, Hunan, Guangdong, Guizhou, and Guangxi) randomly selected from the 16 provinces in this region (Fig. 1). The number of the trees sampled in each province was determined depending on the provincial area, stand origin, and diameter at breast height (D) class in proportion to its contribution to the total stocking volume of the species in the entire region. The number of the sample trees allocated to natural and planted forests was also

determined on a pro rata basis with 77 trees selected from the former and 73 trees from the latter. The sample trees were distributed evenly among D classes of 2, 4, 6, 8, 12, 16, 20, 26, 32, and ≥ 38 cm (i.e., 15 trees for each D class, except for the 26 and 32 cm classes in which there were 14 and 16 trees, respectively). Additionally, the sample trees in each D class were classified, as evenly as possible, into three to five height classes, with three to five trees in each height class. The sample trees represented the characteristics of forests in each secondary climate zone in subtropical China. Their percentages among the D classes and climate zones (north-, middle-, and south-subtropical zones and the entire subtropical zone) are shown in Fig. 2a, b, respectively.

The diameter at breast height (D) was measured on each standing sample tree in the field. After a tree was felled, the total tree length, the length of its living crown, and the diameter at ground level were also measured. The fresh masses of stem wood, stem bark, branch, and foliage were weighted, and all the subsamples (stem sections of a sample tree) were selected and weighed in the field for water content analysis. After being taken to the laboratory, all the subsamples were oven-dried at 85 °C until a constant dry weight was reached. Using the dry weight to fresh weight ratio, the biomass of each component (wood, bark, branch, and foliage) was calculated and added together to get the total AGB of each tree. The relationship of AGB with D for each climate zone with stand origin (natural and planted) is shown in Fig. 2c.

The complete biomass dataset was randomly divided into two groups: 120 trees for model calibration (calibration dataset) and the other 30 trees for model validation (validation dataset). The data and relevant stand characteristics for the calibration and validation datasets are summarized in Table 1.

2.2 Climate data

Climate data for the subtropical China were obtained from the WorldClim database (<http://www.worldclim.org>) with 18 candidate bioclimatic variables (Table OR1 in supplementary file) at a 1 km \times 1 km spatial resolution (Hijmans et al. 2005). The growing season was determined using temperature-based rules, starting when the mean daily temperature was greater than 6.0 °C for five consecutive days, and ending when the average temperature fell below 10.5 °C after the first day of August (Feng et al. 2011). The observations of the variables were averaged for the period 1950–2000.

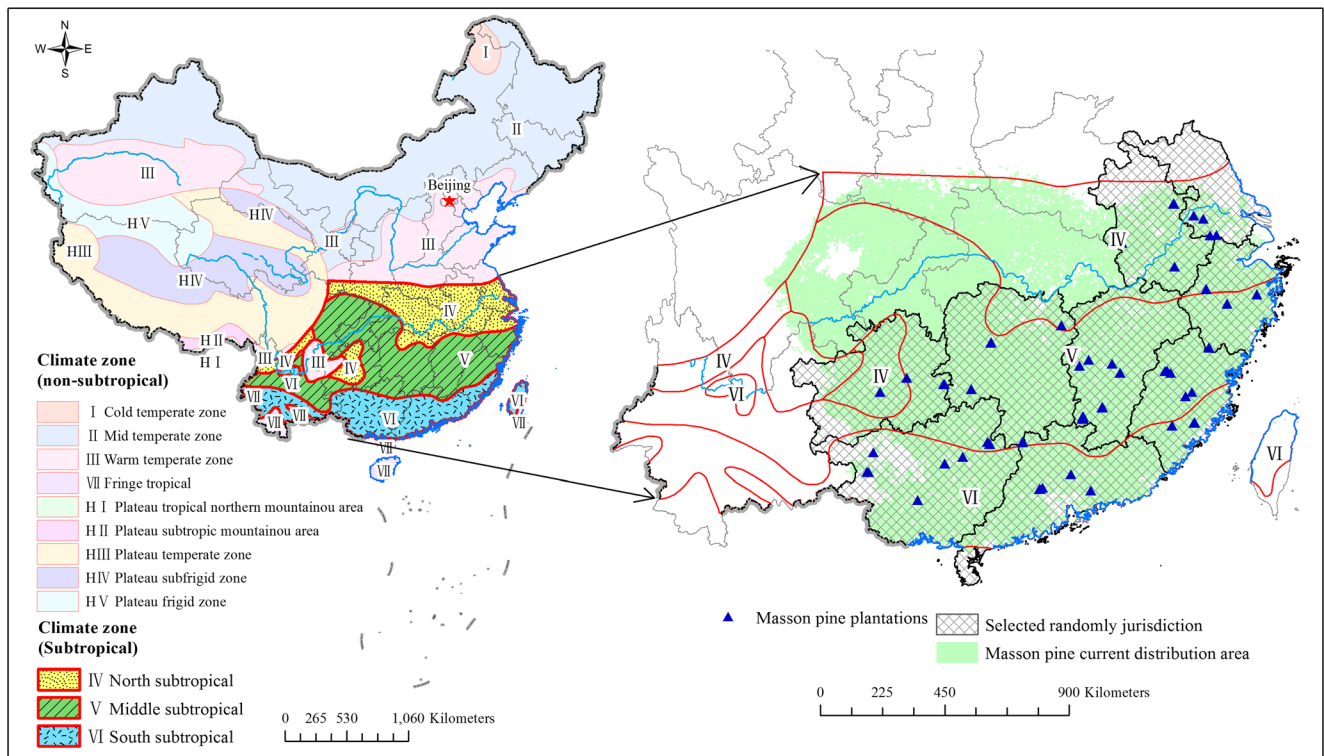


Fig. 1 Distribution of Masson pine sample sites across the subtropical China. The distribution of Masson pine forests in China can be divided into four subregions, namely, the north-subtropical, middle-subtropical,

south-subtropical, and other climate zones (such as mid-temperate zone). However, the fourth subregion had less than 1% of the stand volume of Masson pine in China and was thus not included in this study (SFA 2012)

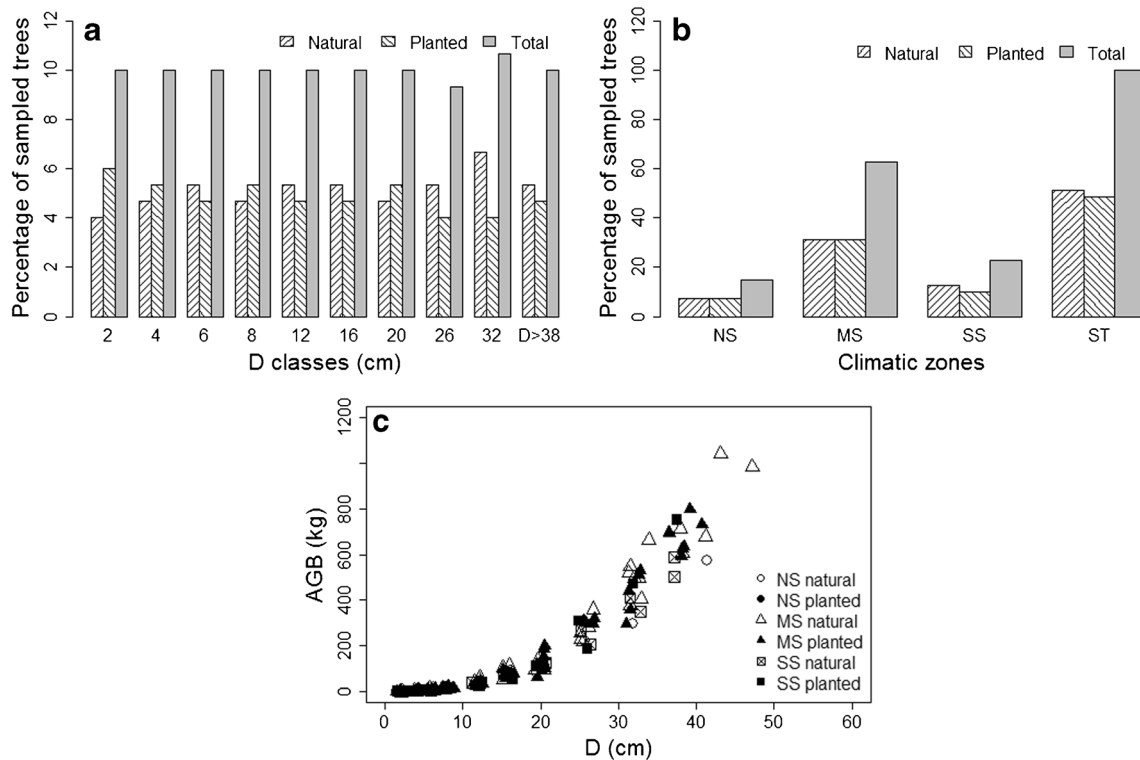


Fig. 2 The percentages of the sampled trees falling. **a** Diameter at breast height (*D*, cm) classes. **b** Climate zones including the north-, middle-, and south-subtropical zones and the entire subtropical zone (denoted by *NS*,

MS, *SS*, and *ST*, respectively). **c** The relationship of aboveground biomass (*AGB*, kg) with *D* for different stand origins with natural and planted forests

Table 1 Summary statistics of diameter at breast height (D) and total aboveground biomass (AGB) of Masson pine by origins (natural and planted forests) and secondary climate zones for model calibration and validation datasets used in this study (N denotes the total number of observations, and SD denotes the standard deviation)

Variable	Calibration data					Validation data				
	N	Mean	SD	Min	Max	N	Mean	SD	Min	Max
Natural										
D (cm)	62	17.38	12.24	2.00	43.10	15	18.45	13.61	2.00	47.20
AGB (kg)	62	178.16	233.07	0.54	1039.14	15	213.35	295.64	0.59	981.78
Planted										
D (cm)	58	15.48	11.36	1.50	40.70	15	15.48	13.62	2.10	39.20
AGB (kg)	58	146.42	209.30	0.32	751.94	15	175.08	272.75	0.75	797.05
North-subtropical										
D (cm)	19	11.84	11.01	1.60	41.40	3	10.07	5.42	5.90	16.20
AGB (kg)	19	77.14	144.87	0.54	575.09	3	34.37	43.09	4.08	83.71
Middle-subtropical										
D (cm)	73	17.82	11.62	1.50	43.10	21	19.13	15.04	2.00	47.20
AGB (kg)	73	188.48	237.14	0.38	1039.14	21	250.65	315.57	0.59	981.78
South-subtropical										
D (cm)	28	16.06	12.41	1.70	37.50	6	12.83	8.18	5.90	25.20
AGB (kg)	28	154.07	212.94	0.32	751.94	6	76.62	104.13	4.91	268.05
Total										
D (cm)	120	16.47	11.81	1.50	43.10	30	16.96	13.47	2.00	47.20
AGB (kg)	120	162.82	221.54	0.32	1039.14	30	194.22	280.16	0.59	981.78

2.3 Climate-sensitive aboveground biomass model

Allometric models are most commonly used to estimate tree biomass by relating tree biomass to easily measured variables (e.g., D) (Zeng et al. 2011; Chave 2014). Zeng et al. (2011) developed an allometric model (model 1) for Masson pine AGB estimation based on D , coupled with a dummy variable approach to address tree origins. They reported that their AGB model provided higher predictive accuracy than other Masson pine AGB models tested. We selected model (1) as the base function to develop our AGB models in this study.

$$\text{AGB} = \exp(\beta_0 + L \times S)D^{\beta_1} + \varepsilon, \quad (1)$$

where AGB is the aboveground biomass (kg); D is the diameter at breast height (cm); S is a dummy variable representing the stand origins, which takes a value 1 for plantations and 0 for natural forests; β_0 , β_1 , and L are model coefficients; and ε is the error term.

The allometric scaling theory suggests the existence of a universal power-law relationship between tree biomass and D with a fixed allometric power (e.g., close to 8/3) which has been verified by numerous studies for various tree species in geographically and ecologically different regions around the

world (West et al. 1997, 1999; Vieilledent et al. 2012; Anitha et al. 2015). Therefore, in this study, it was assumed that climate would only affect the allometric scale, that is, intercept term β_0 in model (1) in the allometric model. The parameter was modified by the subtropical zone-related dummy variables and climatic variables. Firstly, we introduced two additional dummy variables, T_1 and T_2 , to represent differences among three secondary climate zones in the subtropical China: $T_1 = 1$ and $T_2 = 0$ denoted the south-subtropical zone; $T_1 = 0$ and $T_2 = 1$ the middle-subtropical zone; and $T_1 = 0$ and $T_2 = 0$ the north-subtropical zone. Then, the AGB model took the following form

$$\text{AGB} = \exp(\beta_0 + L \times S + k_1 T_1 + k_2 T_2)D^{\beta_1} + \varepsilon, \quad (2)$$

where k_1 and k_2 were the model coefficients, and the other variables were defined as above for model (1).

To explicitly include climatic variables, we modified model (2) by incorporating the temperature- and precipitation-related variables described previously. Model (2) was first fitted to the calibration dataset for Masson pine using generalized least-squares (GLS) in R (R Development Core Team 2011). Then, a combination of temperature- and precipitation-related variables was introduced to the model and evaluated for their contributions to model improvement. The climate variables that

Table 2 Parameter estimates and fit statistics (standard error (SE), Akaike's Information Criterion (AIC), logarithm likelihood (log-likelihood)) of models (2) and (3) using generalized least-squares, and model (7) using a nonlinear mixed-effects model approach for Masson

pine data from the subtropical China (β_0 – β_5 , k_1 , k_2 , and L are the model coefficients; σ^2 is the scaling factor of the error dispersion; and u_3 and u_5 represent the random-effect coefficients)

Parameters	Model (2)		Model (3)		Model (7)	
	Estimates	SE	Estimates	SE	Estimates	SE
β_0	-2.9435	0.2809	-2.7271	1.3385	-1.4544	2.2098
k_1	0.2876	0.0680	-0.0285	0.1079	-0.0175	0.1237
k_2	0.3821	0.0632	0.0077	0.1069	0.0767	0.0819
L	-0.0031	0.0292	0.0240	0.0272	-0.0285	0.0471
β_1	2.5037	0.0760	2.4715	0.0683	2.3697	0.0296
β_2			-0.9078	0.4667	-0.7053	0.5677
β_3			1.1108	0.4566	0.7337	0.4842
β_4			-1.1991	0.3104	-0.3961	0.3679
β_5			1.0760	0.4030	0.4973	0.5346
Variance components						
σ^2	0.5792		0.4380		0.0450	
var(u_3)					0.0001	
var(u_5)					0.00003	
cov(u_3 , u_5)					0.00001	
Autocorrelation structure						
ρ					0.1412	
Weight (power of D)						
γ					2.1463	
AIC	1241.4790		1214.1370		889.7032	
Log-likelihood	-614.7396		-597.0685		-429.8516	

were both significant ($\alpha = 0.05$) and improved model fitting were selected. During the variable selection, we prioritized choosing a set of variables with minimum multicollinearity assessed using the variance inflation factor.

The evaluation of combining the temperature- and precipitation-related variables indicated that the total growing season precipitation (TGSP), mean growing season temperature (MGST), mean temperature of wettest quarter (MTWQ), and precipitation of wettest quarter (PWQ) were significant in explaining variations in the AGB of Masson pine ($p < 0.05$) (Table 2). The spatial distributions of the four climate variables for the period 1950–2000 are displayed in Fig. 3. After the variables were involved, the model (hereafter referred to as the full model) was as follows

$$AGB = \exp(\beta_0 + L \times S + k_1 T_1 + k_2 T_2) D^{\beta_1} TGSP^{\beta_4} MGST^{\beta_3} MTWQ^{\beta_4} PWQ^{\beta_5} + \varepsilon \quad (3)$$

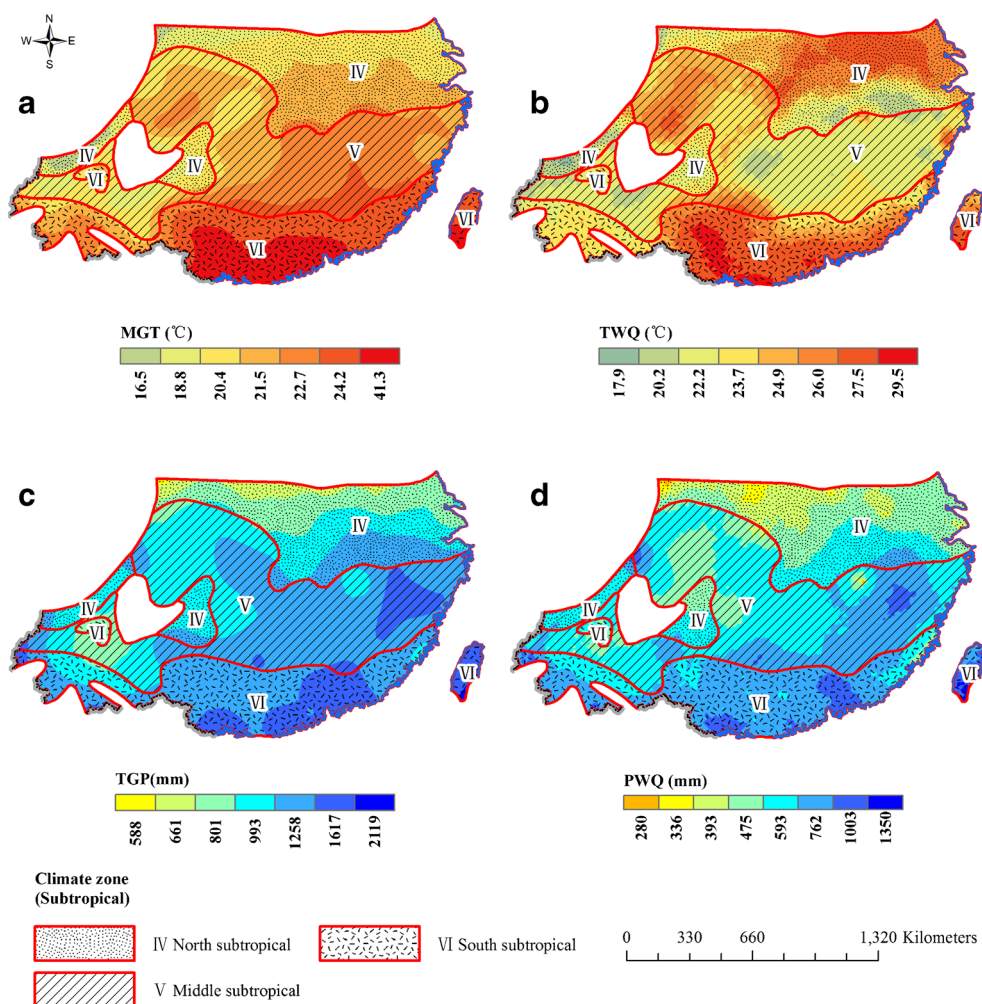
where β_0 – β_5 , k_1 , L , and L are the model coefficients, and the other variables were as defined previously.

When the long-term averages of the four climate variables were included, Akaike's information criterion (AIC) and the logarithm of the likelihood function (log-likelihood) indicated that the model fitness was improved (Table 2).

2.4 Model calibration and evaluation

When the relationship of tree AGB with stand or tree variables are modeled, the measurements of tree variables are often collected from trees growing in sample plots located in different stands with different regions. This hierarchical structure (i.e., trees within regions) can lead to dependence among the measurements, as the observations from the same sampling unit (i.e., each province in this study) are likely to be significantly correlated (i.e., autocorrelation likely exists) (Pinheiro and Bates 2000). To address the problem of autocorrelation, nonlinear mixed-effects modeling (e.g., Fu et al. 2014a, 2014b), recognition of the correlation structure (e.g., Diéguez-Aranda et al. 2006), or both (e.g., Trincado and

Fig. 3 Surface maps of average mean growing season temperature (*MGT*) (a), average mean temperature of wettest quarter (*TWQ*) (b), average total growing season precipitation (*TGP*) (c), and average precipitation of wettest quarter (*PWQ*) (d) for the period from 1950 to 2000 in three secondary climate zones (north-, middle-, and south-subtropical) in the subtropical China



Burkhart 2006) have been used. We chose to use the nonlinear mixed-effects modeling approach provided in the S-Plus *nlme* function (Pinheiro and Bates 2000) to develop the full model. Random-effect parameters were added sequentially, starting with one coefficient as a random effect.

The residuals of the AGB prediction model with random grouping (i.e., province) effects were analyzed for possible autocorrelation or heteroscedasticity. To account for within-province autocorrelation and heteroscedasticity in the variance-covariance matrix (\mathbf{R}_i) of the error term (ε_i) (the subscript i denotes the i^{th} province, $i = 1, \dots, 9$), the approach of Davidian and Giltinan (1995) was used

$$\mathbf{R}_i = \sigma^2 \mathbf{G}_i^{0.5} \mathbf{T}_i \mathbf{G}_i^{0.5} \quad (4)$$

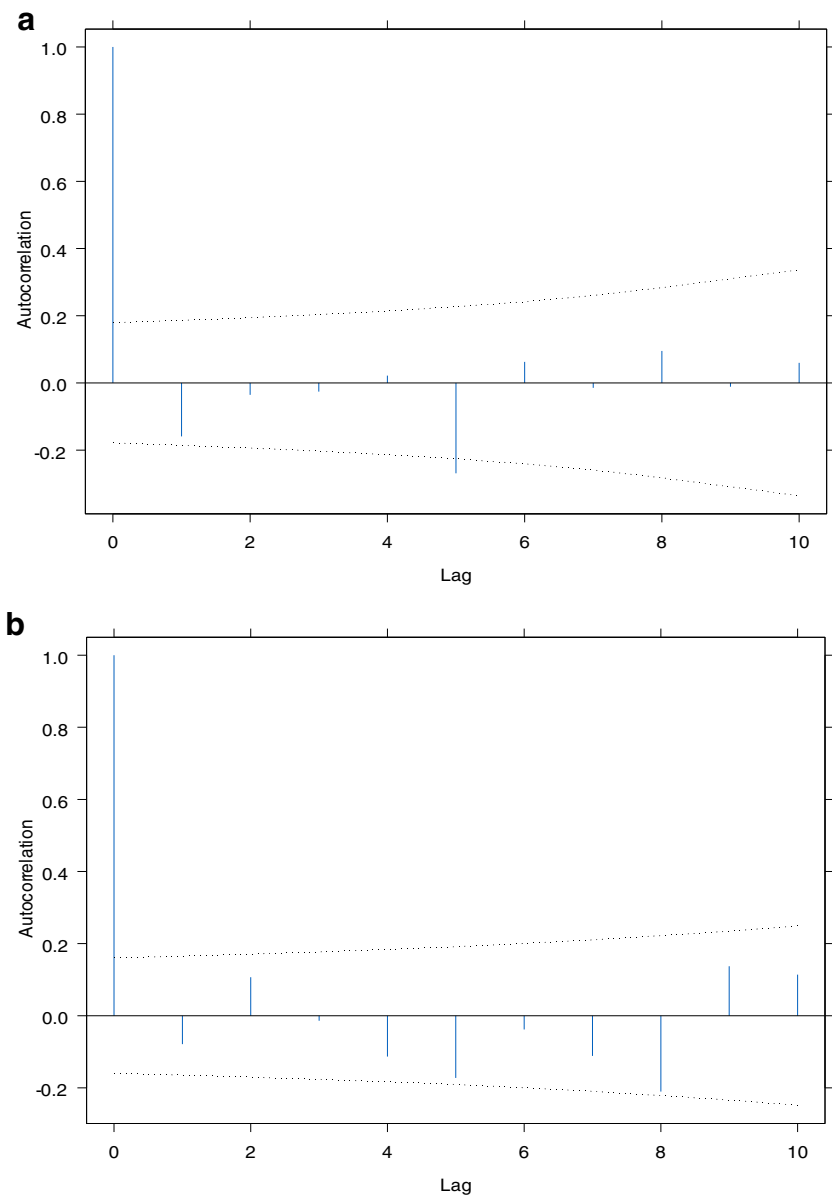
where σ^2 is a scaling factor of the error dispersion, which is equal to the residual variance of the estimated model; \mathbf{G}_i is a

$n_i \times n_i$ diagonal matrix of the within-province variances; and \mathbf{T}_i is a $n_i \times n_i$ matrix showing the within-province autocorrelation structure of the errors.

The empirical autocorrelation function (ACF) for model (3) indicated that the autocorrelation was significant ($\alpha = 0.05$) among the residuals within the provinces (Fig. 4a). As a result, the final model was estimated as a nonlinear mixed-effects model, accounting for the within-province autocorrelation. The autocorrelation structure autoregressive process of order one (AR (1)) for matrix \mathbf{T}_i was applied in this study.

Preliminary analysis for fitting model (3) showed clear evidence of heteroscedasticity in our datasets (Fig. 5a). Some heteroscedasticity remained even after incorporating the autocorrelation structure in the nonlinear mixed-effects model. The heteroscedasticity was removed by specifying a within-province variance function (Pinheiro and Bates 2000). Three common variance functions, e.g., the exponential function [$\text{var}(\varepsilon_{ij}) = \sigma^2 \exp(2\gamma D_{ij})$], the power function

Fig. 4 Empirical autocorrelation function (ACF) corresponding to the standardized residuals of the Masson pine biomass data for model (3) (a) and model (7) (b) with an autoregressive process of order one (AR (1)) autocorrelation structure ($\alpha = 0.05$)



$[\text{var}(\varepsilon_{ij}) = \sigma^2 |D_{ij}|^{2\gamma}]$, and the constant plus power function $[\text{var}(\varepsilon_{ij}) = \sigma^2 (\gamma_1 + D_{ij}^{\gamma_2})^2]$ were evaluated, where D_{ij} is the diameter at breast height of the j th tree in the i th province, and $\gamma, \gamma_1, \gamma_2$ and γ_3 are estimated parameters (Pinheiro and Bates 2000). The function that gave the best fit (e.g., the smaller AIC) and provided the smallest residuals with better behavior was selected. The coefficients of the final model were estimated using the restricted maximum likelihood method (Pinheiro and Bates 2000).

We assessed the accuracy of our AGB model using both the calibration and validation datasets in Table 1 based on the mean bias \bar{e} (observed AGB – predicted AGB), the variance of residuals (v), and the root mean square error (RMSE)

($\text{RMSE} = \sqrt{\bar{e}^2 + v}$). RMSE, which combines the mean bias and the variation of the residuals, was used as the primary criterion for AGB model evaluations.

2.5 Potential effects of changing allometry on tree biomass estimates under future climate scenarios

Model (3) was applied to predict the future AGB of Masson pine trees with the same dataset (Table 1) on each sampled site under future climate scenarios in the three secondary climate zones of subtropical China. The relative difference (R_{AGB}) in AGB between current and future climate scenarios was derived (Eq. (5)). For avoiding the uncertainties from D predictions in the future

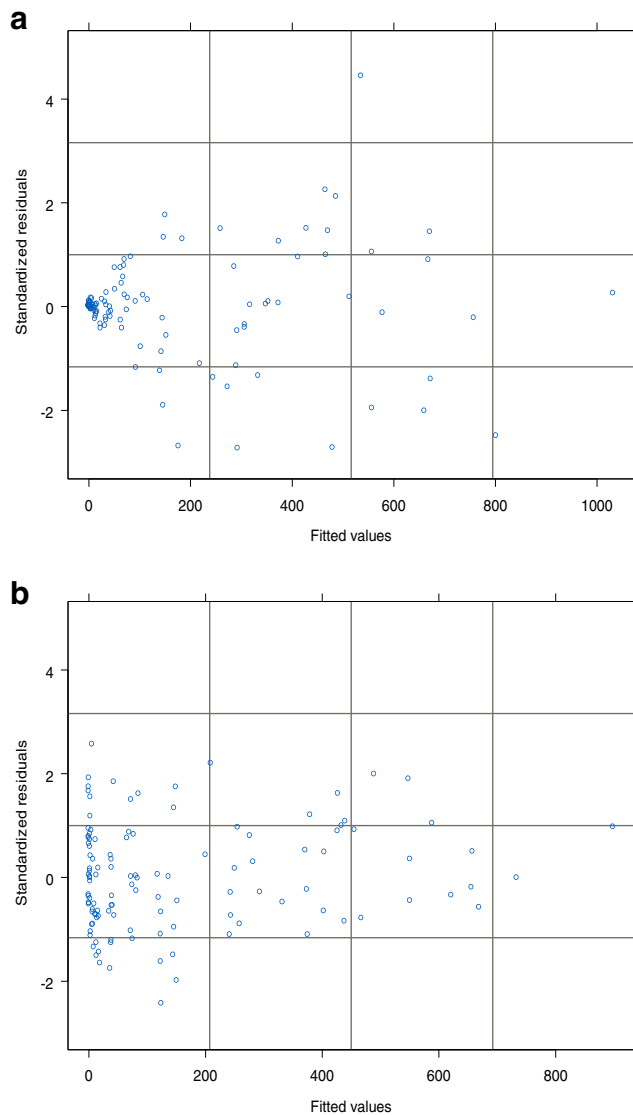


Fig. 5 Distribution of standard residuals for model (3) (a) and model (7) (b) predictions of the aboveground biomass of Masson pine trees in the subtropical China

and making reasonable comparisons, the D values (that is, observed values) of the sample trees were considered constant and applied for estimating the future AGB of Masson pine tree under each climate change scenario. Thus, the effects of different climate scenarios in the future on AGB variation in the different climate zones (e.g., north-, middle-, and south-subtropical zones and the entire subtropical zone) were evaluated under the constraint of the same D . We used the latest climate change scenarios of the fifth Assessment Report from the IPCC based on a down-scaled global climate model (GCM) applied in three

Table 3 Performance assessment of nonlinear mixed-effects model (6) using the Masson pine biomass data with autocorrelation structure autoregressive process of order one (AR (1)) and three variance functions (exponential function, power function, and constant plus power function), respectively (AIC denotes the Akaike information criterion; Log-likelihood denotes the logarithm likelihood; and variance function 1 means that the variances are homogeneous)

Type	AIC	Log-likelihood
Autocorrelation structure AR (1)	889.7032	-429.8516
Variance functions		
Exponential function	956.9301	-464.465
Power function	889.5079	-430.7540
Constant plus power function	890.9463	-430.7732
1	1193.8790	-583.9392

representative concentration pathways (RCPs): RCP 2.6, RCP 4.5, and RCP 8.5 (IPCC 2013). The scenarios indicated low, medium, and high concentrations of greenhouse gases and predictive radiative forcing. The GCM was Beijing Climate Center Climate System Model (BCC-CSM1-1) developed in China. The future climate data at a spatial resolution of 1 km × 1 km for the two periods 2041–2060 and 2061–2080 were downloaded from the WorldClim dataset (Hijmans et al. 2005).

$$R_{AGB} = \frac{AGB_{f|D} - AGB_{c|D}}{AGB_{c|D}} \times 100\% \tag{5}$$

where R_{AGB} is the relative difference in AGB between current and future climate scenarios, and $AGB_{c|D}$ and $AGB_{f|D}$ are the observed and estimated AGB with the same D under the current and future scenarios, respectively.

3 Results

3.1 Climate-sensitive aboveground biomass model

Considering nine model parameters ($\beta_0 - \beta_5$, L , k_1 , and k_2) involved, a total of 512 different combinations of the random effects at the province level were obtained for the full model (model (3)). Among these nonlinear mixed-effects model alternatives, only 45 models converged with meaningful parameter estimates. The following nonlinear mixed-effects AGB model (6) showed the smallest AIC (AIC = 1193.88) and the largest log-likelihood (log-likelihood = -583.94) among the models that converged

$$AGB_{ij} = \left[\exp(\beta_0 + L \times S + k_1 T_1 + k_2 T_2) TGSP_{ij}^{\beta_2} MGST_{ij}^{(\beta_3 + u_{3i})} MTWQ_{ij}^{\beta_4} PWQ_{ij}^{(\beta_5 + u_{5i})} \right] D_{ij}^{\beta_1} + \varepsilon_{ij} \tag{6}$$

where AGB_{ij} is the aboveground biomass of the j th tree in the i th province ($j = 1, \dots, n_i; i = 1, \dots, 9$); n_i is the total observations in the i th province; $TGSP_{ij}$, $MGST_{ij}$, $MTWQ_{ij}$, and PWQ_{ij} are the total growing season precipitation (mm), mean growing season temperature ($^{\circ}C$), mean temperature of wettest quarter ($^{\circ}C$), and precipitation of wettest quarter (mm) for the j th tree in the i th province, respectively; and ε_{ij} is the estimation error for the j th tree in the i th province. u_{3i} and u_{5i} represent the random-effect coefficients generated for the i th province, and other variables and parameters were as defined previously.

The evaluation indices of autocorrelation structure AR (1) and three variance functions applied to model (6), respectively, are listed in Table 3. For comparison purpose, the

$$\begin{cases} AGB_{ij} = \left[\exp(\beta_0 + L \times S + k_1 T_1 + k_2 T_2) TGSP_{ij}^{\beta_2} MGST_{ij}^{(\beta_3 + u_{3i})} MTWQ_{ij}^{\beta_4} PWQ_{ij}^{(\beta_5 + u_{5i})} \right] D_{ij}^{\beta_1} + \varepsilon_{ij} \\ \varepsilon_i = (\varepsilon_{i1}, \dots, \varepsilon_{in_i})^T \sim N(0, \mathbf{R}_i = \sigma^2 \mathbf{G}_i^{0.5} \mathbf{\Gamma}_i \mathbf{G}_i^{0.5}) \\ \mathbf{G}_i = \text{diag}(|D_{i1}|^{2\gamma}, \dots, |D_{in_i}|^{2\gamma}) \\ \mathbf{\Gamma}_i \sim AR(1) \end{cases} \quad (7)$$

where the superscript T denotes the matrix transpose operation, and the other variables and coefficients were as defined previously.

3.2 Estimation of parameters

All estimated coefficients of model (7) were significant ($\alpha = 0.05$) (Table 2). Coefficients β_3 and β_5 were positive, which implied that a warmer growing season or an increase in PWQ favored AGB accumulation. Coefficients β_2 and β_4 were negative, suggesting that excessive rain during the growing season, or high mean temperature in the wettest quarter, would decrease AGB. The negative coefficient for L implied that the AGB of Masson pine tree of natural origin exceeded the AGB of planted trees. Coefficient k_1 was negative and k_2 was positive, which showed that the AGB values of trees with the same S , D , $TGSP$, $MGST$, $MTWQ$, and PWQ from the middle-subtropical zone were the highest, followed by those in the north- and south-subtropical zones. The positive sign of ρ indicated that the model error terms were positively correlated with prediction residuals.

3.3 Model evaluation

Models (3) and (7) were further evaluated using both the calibration and validation datasets. The prediction accuracy of model (7) ($\bar{e} = 5.18$ kg, $v = 43.14$, $RMSE = 43.45$ kg for the calibration dataset; $\bar{e} = 1.08$ kg, $v = 45.75$, $RMSE = 45.76$ kg for the validation dataset) was much higher than that of model (3) (\bar{e}

assessment results of model (6), in which the variances of error terms were assumed to be homogeneous (i.e., when \mathbf{G}_i and $\mathbf{\Gamma}_i$ are identity matrices), were also given (Table 3). The AIC and log-likelihood of model (6) with AR (1) were 25.48% smaller and 26.39% larger than those of model (6) with homogeneous assumption, respectively, which indicated that the within-province autocorrelation in the Masson pine biomass data could be successfully removed by the AR (1). The power variance function demonstrated the best performance among the variance functions we examined. Therefore, model (6) was fitted with an AR (1) autocorrelation structure and a power variance function of D . The final non-linear mixed-effects AGB model was as follows

$= 5.30$ kg, $v = 52.56$, $RMSE = 52.82$ kg for the calibration dataset; $\bar{e} = 7.87$ kg, $v = 46.13$, $RMSE = 46.80$ kg for the validation dataset). Especially, the $RMSE$ of model (7) based on the calibration dataset was 17.7% smaller than that of model (3), and the mean bias of model (7) based on the validation dataset was 86.3% lower than that of model (3). The residuals from the calibration data for model (7) displayed no discernable patterns, indicating that heteroscedasticity was effectively accounted for by the power variance function (Fig. 5b). Similar residual patterns were found for the validation dataset, and the error structures did not show any signs of autocorrelation (Fig. 4b).

3.4 Potential effects of changing allometry on tree biomass estimates under future climate scenarios

Figure 6 shows the box plots and distributions of the R_{AGB} in predicted AGB of Masson pine trees between the current and climate scenarios—RCP 2.6, RCP 4.5, and RCP 8.5—during the periods of 2041–2060 and 2061–2080 for the different climate zones (north-, middle-, and south-subtropical zones and the entire subtropical zone) and tree origins (natural and planted). The results indicated that both the increase and the decrease of AGB occurred in the subtropical China. Generally, the predicted values of tree AGB in three secondary climate zones under each climate scenario for the periods of 2041–2060 and 2061–2080 increased with the increased regional annual mean temperature and mean total precipitation. The differences of the increases in AGB among the climate zones for the same time period under three climate scenarios

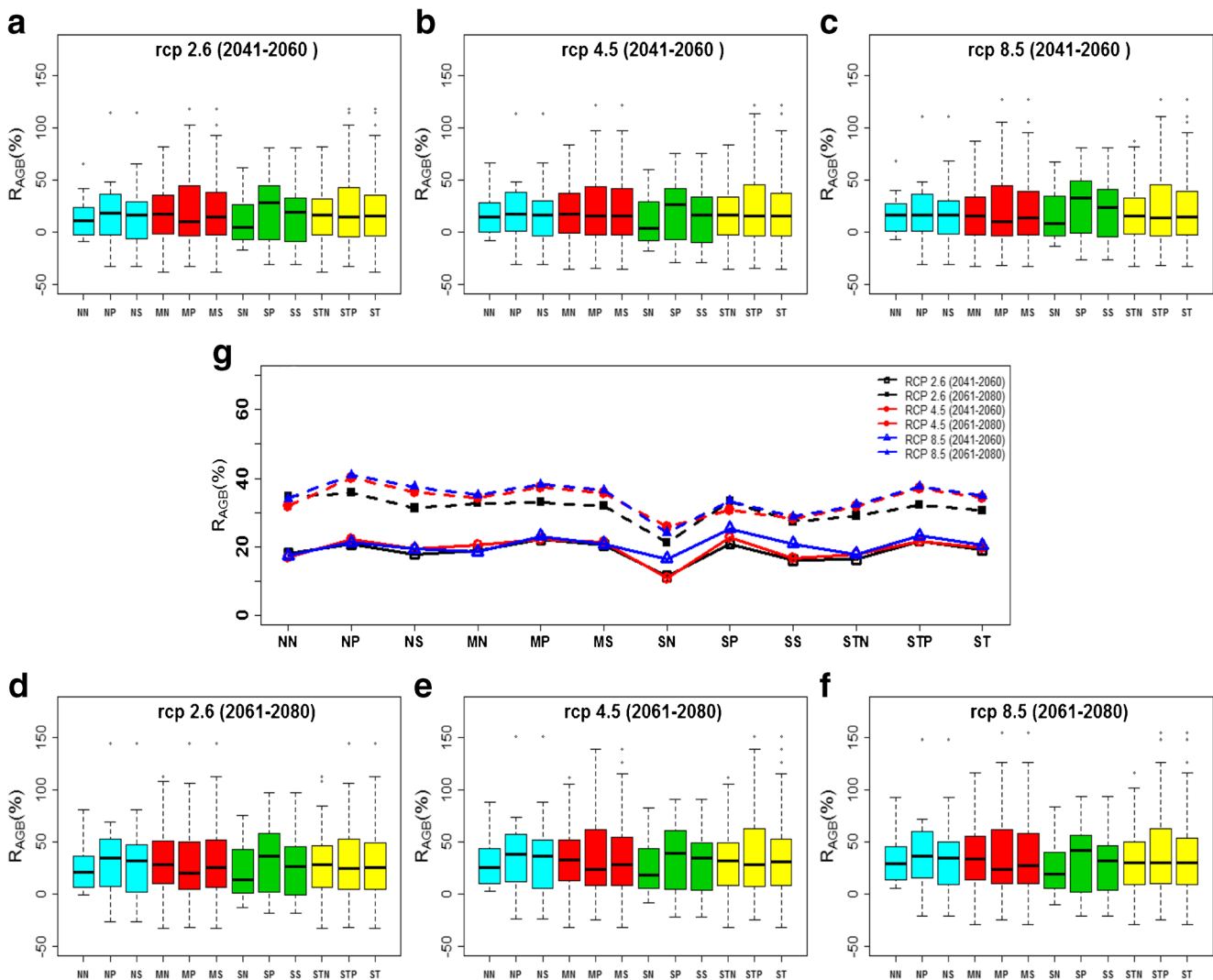


Fig. 6 The relative differences (R_{AGB}) in predicted aboveground biomass (AGB) of Masson pine trees between the current and future climate scenarios: RCP 2.6, RCP 4.5, and RCP 8.5 during the periods of 2041–2060 and 2061–2080 for three different secondary climate zones and the entire subtropical zone in China: **a–f** the box plots of R_{AGB} and **g** the distributions of the mean values of R_{AGB} ; NN, MN, SN, and STN

were substantially large (Fig. 6). The estimated average of AGB under each climate scenario in the secondary climate zones was the highest for the period 2061–2080, followed by the period 2041–2060, and the period 1950–2000 (Fig. 6g).

With other variables (e.g., D , secondary climate zone, and climate scenario) being equal, the estimated AGB of planted trees was higher than that of natural trees for the periods 2041–2060 and 2061–2080 (Fig. 6). The greatest increase in AGB was predicted for the planted trees in the north-subtropical zone for the period 2061–2080 under the RCP 8.5 climate scenario (overall increasing by $41.0 \pm 44.6\%$ (mean \pm standard deviation) over the current AGB). For the time periods 2041–2060 and 2061–2080, Fig. 6g also shows that overall, the estimated average AGB under each secondary climate zone was the highest for the RCP 8.5, followed by the

denote trees of natural origin in the north-, middle-, and south-subtropical zones and the entire subtropical zone, respectively; NP, MP, SP, and STP denote trees of plantations in the north-, middle-, and south-subtropical zones and the entire subtropical zone, respectively; NS, MS, SS, and ST denote all trees in the north-, middle- and south-subtropical zones and the entire subtropical zone

RCP 4.5, and the lowest for the RCP 2.6. The tree AGB of Masson pine in the entirety of the subtropical China was predicted to increase from the current levels under the RCP 2.6, RCP 4.5, and RCP 8.5 climate scenarios by 19.0 ± 30.6 , 19.6 ± 30.4 , and $20.5 \pm 31.0\%$ for the period 2041–2060 (Fig. 6a–c); by 30.6 ± 33.6 , 34.5 ± 34.7 , and $34.8 \pm 35.0\%$ for the period 2061–2080 (Fig. 6d–f), respectively; and by 24.8 ± 32.7 , 27.0 ± 33.4 , and $27.7 \pm 33.8\%$ for the whole periods 2041–2080, respectively.

4 Discussion

Tree growth and AGB accumulation differ with different climate, site conditions, and management practices (Landsberg

and Sands 2011). Climate variables need to be incorporated into AGB models of trees if these models are utilized to directly account for the impact of changing climate on tree biomass accumulation. This can be accomplished, at least in the composite model form (e.g., Wykoff 1990), by modifying the intercept term (i.e., β_0) in biomass models such as model (2). In this study, it was assumed that climatic variables only influence the allometric scale of AGB equation with the observation of a universal allometric power (West et al. 1997, 1999; Vieilledent et al. 2012; Anitha et al. 2015) although there were some exceptions (Chave et al. 2005). The space–time replacement approach (Pickett 1989) was applied in this study to develop the climate-sensitive AGB model. Results showed that the inclusion of climate variability in the allometric scale of the biomass equation yielded improved estimation and the scale power was close to $8/3$ which supported our assumption.

It was found that a combination of the temperature- and precipitation-related variables (TGSP, MTWQ, MGST, and PWQ) was significant in explaining the variation in AGB of Masson pine trees across three secondary climate zones in the subtropical China. The modeling results indicated that more precipitation over the growing season negatively affected the AGB of Masson pine trees. Previous studies indicated that the relationship between mean annual precipitation and tree AGB or net primary productivity, on the most of sites, could be described by a nonlinear, concave-down function (Hsu et al. 2012). The decline in the AGB was likely an indirect effect mediated by the availability of other resources (Hsu et al. 2012; Subedi and Sharma 2013). While additional moisture usually increases tree biomass in dry-to-mesic ecosystems, increasing precipitation in humid ecosystems may reduce tree biomass by decreasing radiation, increasing nutrient leaching, or reducing soil oxygen availability (Schuur 2003; Hsu et al. 2012). Increasing precipitation may also reduce the diffusion of oxygen through water-filled pores, as well as decomposition rates (Subedi and Sharma 2013). In addition, the results that an increase in precipitation led to the decrease of AGB could be also explained by biological traits of Masson pine. For example, Masson pine trees mostly grow well in the ridges of hills or mountains with good light and dry sites (Xia et al. 1996; Zhang et al. 2013).

However, more precipitation during the wettest quarter positively impacted the AGB of Masson pine trees in our model. The wettest period in the subtropical China occurs during the months of May, June, July, and August (falling in the second and third quarters). This period also has the highest yearly temperatures (Yang et al. 2006); thus, it had the highest AGB of Masson pine trees. This is consistent with the findings from the previous study by Zhang et al. (2013).

Understanding seasonal growth patterns of the subtropical coniferous trees is important to explain the relationship between moisture availability and tree AGB. In the case for temperate trees (Tardif et al. 2001), the cumulative annual

biomass growth of Masson pine trees can be described using three distinct phases. In March to April, an initial period of swelling of a stem occurs, indicating the upward passage of water. This is followed by the period (May–November) of active cell division, indicating the “grand period” of growth. During the third period (December to February), tree growth stops, the cells dehydrate, and the cambial tissue rests. Optimal growing conditions during the grand period are very important for Masson pine trees. The positive relationship that we found between increased moisture availability during the “grand period of growth” and AGB was as expected.

It was found that a warmer growing season increased the AGB of Masson pine trees, which is consistent with previous findings (e.g., Gholz 1982). The warming-induced increase in tree productivity may be a direct effect of either increased photosynthesis at higher temperatures, longer growing seasons, or an indirect effect of increased nutrient availability, resulting from increased rates of litter decomposition and N mineralization. Both the direct and indirect effects of the warming could be particularly important in the north-subtropical zone in which both temperature and nutrient limitations are more critical relative to the other two subzones.

The mean temperature during the wettest quarter appeared to be the main factor in limiting AGB of trees (Table 2). It is likely that increasing temperatures increase evaporative demand, causing increased water stress, which may impact tree growth more strongly than the increase in temperature alone (Ciais et al. 2005). The temperature in the subtropical China in the wettest quarter is quite high overall, thereby increasing evapotranspiration and reducing the availability of soil moisture to trees. Therefore, trees in this period may have less available moisture for photosynthesis and consequently, lower growth. This is also supported by a simulation study of seasonal photosynthesis in the northwest USA, which indicated that growing-season photosynthesis was greatly restricted by drought (Emmingham and Waring 1977).

The results showed that there was the highest AGB of Masson pine trees in the middle-subtropical zone, followed by the north- and south-subtropical zones (Table 2; Fig. 6). This ranking is similar to that found by Yang et al. (2006). In the middle-subtropical zone, the number of days without frost, or with a mean daily temperature ≥ 10 °C, is more than 230 days, the corresponding accumulated temperature is above 5000 °C, and the range of annual mean precipitation is 1000–2500 mm. These are favorable conditions for tree growing (Yang et al. 2006). In the north-subtropical zone, although there is abundant precipitation (the range of annual mean precipitation is 800–1800 mm), the temperature slightly lower than that in the middle-subtropical zone reduces the potential for tree growing (Yang et al. 2006). In the south-subtropical zone, precipitation is quite close to the middle-subtropical zone, but the mean temperature is higher. Therefore, the evaporation in this subregion is higher than that

in the middle-subtropical zone, which, in turn, decreases the biomass of trees (Yang et al. 2006).

When the effects of future climate change on AGB using the climate-sensitive AGB model developed in the study were analyzed, ideally, the D growth of sample trees should be predicted related to climate change (Zeng et al. 2017). The climatic impacts on AGB of each sample tree under climate change scenarios should be then analyzed. However, in practice, in addition to climatic zones and climatic-related variables, other factors, such as tree height, crown size, development stage, forest types, site characteristics, stand structure, stand density (including tree recruitment and death), and competition, might significantly contribute to the variation in the individual tree D (Subedi and Sharma 2013; Ashraf et al. 2015). Some of these factors interact with each other and vary with future climate change, which would result in the complexity and very large uncertainty in future AGB estimates of sample trees for this method when D growth is considered. In addition, a fixed allometric power in the power-law relationship between tree biomass and D was commonly assumed in the biomass modeling, which was further verified in this study (Table 2). Relative to traditional biomass allometric equations [e.g., model (1)], the developed climate-sensitive AGB model [model (7)] accounted for potential climate-biomass relationships effectively (Table 2) by parameterization of allometric scale. Thus, the effects of different future climate scenarios on AGB variation were evaluated under the constraint of the same D in this study. Additionally, the effects of wood density on AGB were also observed (Zhang et al. 2012; Deng et al. 2014) and used for improving biomass estimation (Chave et al. 2014). Unfortunately, we did not measure the wood density of each sample tree in the data collection. But, the effect of wood density may be indirectly reflected by climatic factors and random effects.

According to the climate-sensitive AGB model developed in the study, we estimated the AGB of Masson pine trees in the subtropical China under future climate changes and found that the AGB would increase by 10.91 to 41.01% under the RCP 2.6, RCP 4.5, and RCP 8.5 climate scenarios for the period 2041–2080 (Fig. 6). Therefore, traditional biomass equation without climate variables will produce some bias in biomass and carbon estimation as expected. The magnitude of the effects of climate change on AGB revealed the variation and uncertainty of AGB predictions with stand origin, climate zones, and climate scenarios. Our study also showed clear positive effects of future climate change on AGB of Masson pine trees. Medlyn et al. (2011) summarized the uncertainties about the effects of climate change in terms of direction and magnitude on forest productivity due to the differences in model strategy, climate variables, climate scenarios, tree species, and regions. The uncertainties need to further be addressed when the potential factors affecting tree biomass are investigated in the future.

Acknowledgements The article was funded by the Forestry Public Welfare Scientific Research Project (No. 201504303) and the Chinese National Natural Science Foundation (Nos. 31470641, 31300534, and 31570628). We appreciate the valuable comments from the chief editor Dr. Erwin Dreyer, Dr. Jean-Michel Leban, the handling editor, and two anonymous reviewers who improved the manuscript.

Compliance with ethical standards The biomass data used in this study were collected according to the protocol of data collection for tree biomass modeling which was drafted by the State Forestry Administration of China (SFA 2014).

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